

Supplementary Materials

A Proof of Theorem 1

A.1 Key Properties

We provide a number of concentration properties under non-uniform sampling. These properties are in parallel to those under uniform sampling used in [1, 3, 12]. More specifically, Lemma 1 is proven in [9], which readily implies Lemma 3. We develop the proofs for other lemmas based on local incoherence, and provide the detailed proofs in Appendix B.

Lemma 1. [9, Lemma 9] Suppose $\mathbb{P}((i, j) \in \Omega_0) = q_{ij}$ for all $i, j \in [n]$. If $q_{ij} \geq C_0(\mu_{0ij}r \log n)/n$ for some sufficiently large constant C_0 and for all $i, j \in [n]$, then with high probability

$$\|\mathcal{P}_T - \mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T\| \leq \frac{1}{2}. \quad (17)$$

Lemma 2. If $\|\mathcal{P}_T - \mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T\| \leq \frac{1}{2}$ and $p_{ij} \geq p_0$ for all $i, j \in [n]$, then

$$(a) \quad \|\mathcal{P}_T \mathcal{R}_{\Omega_0}\| \leq \sqrt{\frac{3}{2p_0}};$$

$$(b) \quad \mathcal{P}_{\Omega_0} \mathcal{P}_T \text{ is injective on } T.$$

Lemma 3. Suppose $\mathbb{P}((i, j) \in \Omega_0) = q_{ij}$ for all $i, j \in [n]$. For a fixed matrix $Z \in T$, if $q_{ij} \geq C_0(\mu_{ij}r \log n)/n$ for some sufficiently large constant C_0 and for all $i, j \in [n]$, then with high probability

$$\|Z - \mathcal{P}_T \mathcal{R}_{\Omega_0}(Z)\|_F \leq \frac{1}{2} \|Z\|_F. \quad (18)$$

Lemma 4. Suppose $\mathbb{P}((i, j) \in \Omega_0) = q_{ij}$ for all $i, j \in [n]$. For a fixed matrix $Z \in T$, if $q_{ij} \geq C_0 \sqrt{\mu_{ij}r/n}$ for some sufficiently large constant C_0 and for all $i, j \in [n]$, then with high probability

$$\|(\mathcal{R}_{\Omega_0} - I)Z\| \leq \frac{C}{C_0} \|Z\|_{w(\infty)} \quad (19)$$

for some constant C .

Lemma 5. Suppose $\mathbb{P}((i, j) \in \Omega_0) = q_{ij}$ for all $i, j \in [n]$. Suppose $\beta > 0$ is a scaling factor. For a fixed matrix $Z \in T$, if $q_{ij} \geq C_0 \beta^{-2} \sqrt{\mu_{ij}r/n}$ for some sufficiently large C_0 and for all $i, j \in [n]$, then with high probability

$$\|(\mathcal{P}_T \mathcal{R}_{\Omega_0} - \mathcal{P}_T)Z\|_{w(\infty)} \leq \frac{1}{2} \beta \|Z\|_{w(\infty)}. \quad (20)$$

Lemma 6. Suppose S is the error matrix in the random sign model defined in Section 2.1. Then for any given index (a, b) with $a, b \in [n]$, with high probability

$$|[\mathcal{P}_T \text{sgn}(S)]_{ab}| \leq C \sqrt{\frac{\mu_{ab}r \log n}{n}} \quad (21)$$

for some constant C .

A.2 Proof of Proposition 1 (Dual Certificate Conditions)

Due to the assumption of the proposition, $\Gamma = \Omega^c$ satisfies the conditions required in Lemma 1. Hence, due to Lemmas 1 and 2, we have $\|\mathcal{P}_T \mathcal{R}_\Gamma\| \leq \sqrt{\frac{3}{2p_0}}$ with $p_0 = 1/n^3$ and $\mathcal{P}_\Gamma \mathcal{P}_T$ is injective on T with high probability.

Suppose $\hat{L} = L + H$ and $\hat{S} = S - H$ satisfy

$$\|L + H\|_* + \lambda \|S - H\|_1 \leq \|L\|_* + \lambda \|S\|_1. \quad (22)$$

By the definition of subgradient, we have

$$\|L + H\|_* \geq \|L\|_* + \langle \mathcal{P}_T H, UV^* \rangle + \|\mathcal{P}_{T^\perp} H\|_*$$

where we use the fact that there exists $W \in T^\perp$ and $\|W\| \leq 1$ such that $\|\mathcal{P}_{T^\perp} H\|_* = \langle \mathcal{P}_{T^\perp} H, W \rangle$.

Thus, we have

$$\langle \mathcal{P}_T H, UV^* \rangle + \|\mathcal{P}_{T^\perp} H\|_* \leq \lambda \|S\|_1 - \lambda \|S - H\|_1.$$

Furthermore,

$$\begin{aligned}\|S - H\|_1 &= \|S - \mathcal{P}_\Omega H\|_1 + \|\mathcal{P}_\Gamma H\|_1 \\ &\geq \|S\|_1 + \langle \text{sgn}(S), -H \rangle + \|\mathcal{P}_\Gamma H\|_1.\end{aligned}$$

Combining the last two inequalities, we have

$$\|\mathcal{P}_{T^\perp} H\|_* + \lambda \|\mathcal{P}_\Gamma H\|_1 \leq \langle H, \lambda \text{sgn}(S) - UV^* \rangle.$$

For a matrix Y that obeys the conditions in the Proposition 1, we derive

$$\begin{aligned}\langle H, \lambda \text{sgn}(S) - UV^* \rangle &= \langle H, Y + \lambda \text{sgn}(S) - UV^* \rangle - \langle H, Y \rangle \\ &= \langle \mathcal{P}_T H, \mathcal{P}_T(Y + \lambda \text{sgn}(S) - UV^*) \rangle + \langle \mathcal{P}_{T^\perp} H, \mathcal{P}_{T^\perp}(Y + \lambda \text{sgn}(S)) \rangle \\ &\quad - \langle \mathcal{P}_\Gamma H, \mathcal{P}_\Gamma Y \rangle - \langle \mathcal{P}_\Omega H, \mathcal{P}_\Omega Y \rangle \\ &\leq \frac{\lambda}{n^2} \|\mathcal{P}_T H\|_F + \frac{1}{4} \|\mathcal{P}_{T^\perp} H\|_* + \frac{\lambda}{4} \|\mathcal{P}_\Gamma H\|_1.\end{aligned}$$

Combining the previous two inequalities, we obtain

$$\frac{3}{4} \|\mathcal{P}_{T^\perp} H\|_* + \frac{3}{4} \lambda \|\mathcal{P}_\Gamma H\|_1 \leq \frac{\lambda}{n^2} \|\mathcal{P}_T H\|_F.$$

We next bound $\|\mathcal{P}_T H\|_F$ as follows:

$$\begin{aligned}\|\mathcal{P}_T H\|_F &\leq 2 \|\mathcal{P}_T \mathcal{R}_\Gamma \mathcal{P}_T(H)\|_F \\ &\leq 2 \|\mathcal{P}_T \mathcal{R}_\Gamma \mathcal{P}_{T^\perp}(H)\|_F + 2 \|\mathcal{P}_T \mathcal{R}_\Gamma(H)\|_F \\ &\leq \sqrt{\frac{6}{p_0}} \|\mathcal{P}_{T^\perp}(H)\|_F + \sqrt{\frac{6}{p_0}} \|\mathcal{P}_\Gamma(H)\|_F.\end{aligned}$$

We thus obtain

$$\left(\frac{3}{4} - \frac{\lambda}{n^2} \sqrt{\frac{6}{p_0}}\right) \|\mathcal{P}_{T^\perp}(H)\|_F + \left(\frac{3}{4} \lambda - \frac{\lambda}{n^2} \sqrt{\frac{6}{p_0}}\right) \|\mathcal{P}_\Gamma(H)\|_F \leq 0.$$

The above inequality implies that if $p_0 \geq 1/n^3$, then $\mathcal{P}_{T^\perp} H = \mathcal{P}_\Gamma H = 0$. This further implies $\mathcal{P}_\Gamma \mathcal{P}_T(H) = 0$. Since $\mathcal{P}_\Gamma \mathcal{P}_T$ is injective on T , we have $\mathcal{P}_T H = 0$. Consequently, $H = 0$.

A.3 Dual Certificate Verification

We show that the dual certificate constructed in (13)-(15) satisfies the conditions in Proposition 1.

We first bound $\|Z_0\|_F$, $\|Z_0\|_\infty$ and $\|Z_0\|_{w(\infty)}$. Observe that for an index pair (a, b) , we have

$$|[Z_0]_{ab}| \leq |[UV^*]_{ab}| + \lambda |[P_T \text{sgn}(S)]_{ab}|.$$

Using the fact that $|[UV^*]_{ab}| \leq \sqrt{\frac{\mu_{ab} r}{n^2}}$ and $\lambda = \frac{1}{32\sqrt{n \log n}}$, and applying Lemma 6, we obtain

$$\|Z_0\|_\infty \leq C\sqrt{\mu r}/n. \quad (23)$$

Furthermore,

$$\|Z_0\|_F \leq \|UV^*\|_F + \lambda \|\mathcal{P}_T \text{sgn}(S)\|_F \leq \sqrt{r} + C\sqrt{\mu r} \leq C'\sqrt{\mu r} \quad (24)$$

where we used $\|Z\|_F \leq n\|Z\|_\infty$ for any matrix Z , and

$$\begin{aligned}\|Z_0\|_{w(\infty)} &\leq \|UV^*\|_{w(\infty)} + \lambda \|\mathcal{P}_T \text{sgn}(S)\|_{w(\infty)} \\ &\leq 1 + \max_{a,b} \lambda \frac{|[P_T \text{sgn}(S)]_{ab}|}{w_{ab}} \\ &\leq C',\end{aligned} \quad (25)$$

where we used the definition $w_{ab} = \max\{\sqrt{\mu_{ab} r}/n^2, \epsilon\}$ and $\lambda = \frac{1}{32\sqrt{n \log n}}$. We note that for the sake of convenience, the constants C and C' may be different from line to line.

We further note that Lemma 3 implies

$$\|Z_k\|_F \leq \frac{1}{2} \|Z_{k-1}\|_F \quad (26)$$

with high probability, provided that $q_{ij} \geq C_0(\mu_{ij}r \log n)/n$ for some sufficiently large constant C_0 and for all $i, j \in [n]$.

Lemma 4 implies

$$\|(I - \mathcal{R}_{\Gamma_k})Z_{k-1}\| \leq \frac{C}{C_0} \|Z_{k-1}\|_{w(\infty)} \quad (27)$$

with high probability, provided that $q_{ij} \geq C_0\sqrt{\mu_{ij}r/n}$ for some sufficiently large constant C_0 and for all $i, j \in [n]$.

Lemma 5 implies

$$\|Z_1\|_{w(\infty)} \leq \frac{1}{2\sqrt{\log n}} \|Z_0\|_{w(\infty)} \quad (28)$$

and

$$\|Z_k\|_{w(\infty)} \leq \frac{1}{2} \|Z_{k-1}\|_{w(\infty)} \quad \text{for } k = 2, \dots, l \quad (29)$$

with high probability, provided that $q_{ij} \geq C_0\sqrt{\mu_{ij}r/n}$ for some sufficiently large constant C_0 and for all $i, j \in [n]$.

We are now ready to show that the constructed dual certificate Y obeys the conditions (9)-(12) in Proposition 1. Clearly, Y satisfies $\mathcal{P}_\Omega Y = 0$ given in (9) due to the construction.

In order to show that Y satisfies (12), we derive

$$\begin{aligned} & \|\mathcal{P}_T Y + \mathcal{P}_T(\lambda \operatorname{sgn}(S) - UV^*)\|_F \\ &= \left\| Z_0 - \left(\sum_{k=1}^l \mathcal{P}_T \mathcal{R}_{\Gamma_k} Z_{k-1} \right) \right\|_F \\ &= \left\| (\mathcal{P}_T - \mathcal{P}_T \mathcal{R}_{\Gamma_1}) Z_0 - \left(\sum_{k=2}^l \mathcal{P}_T \mathcal{R}_{\Gamma_k} Z_{k-1} \right) \right\|_F \\ &= \left\| \mathcal{P}_T Z_1 - \left(\sum_{k=1}^l \mathcal{P}_T \mathcal{R}_{\Gamma_k} Z_{k-1} \right) \right\|_F \\ &= \dots \\ &= \|Z_l\|_F \stackrel{(a)}{\leq} \left(\frac{1}{2}\right)^l \cdot \|Z_0\|_F \stackrel{(b)}{\leq} C' \left(\frac{1}{2}\right)^l \sqrt{\mu r} \leq \frac{\lambda}{n^2}, \end{aligned}$$

where (a) follows from (26) and (b) follows from (24).

In order to show that Y satisfies (11), we respectively show that $\|\mathcal{P}_{T^\perp} Y\| \leq \frac{1}{8}$ and $\|\mathcal{P}_{T^\perp}(\lambda \operatorname{sgn}(S))\| \leq \frac{1}{8}$ as follows.

$$\begin{aligned} \|\mathcal{P}_{T^\perp} Y\| &= \left\| \mathcal{P}_{T^\perp} \sum_{k=1}^l \mathcal{R}_{\Gamma_k} Z_{k-1} \right\| \\ &\leq \sum_{k=1}^l \|\mathcal{P}_{T^\perp} \mathcal{R}_{\Gamma_k} Z_{k-1}\| \\ &\stackrel{(a)}{=} \sum_{k=1}^l \|\mathcal{P}_{T^\perp} (\mathcal{R}_{\Gamma_k} Z_{k-1} - Z_{k-1})\| \\ &\leq \sum_{k=1}^l \|\mathcal{R}_{\Gamma_k} Z_{k-1} - Z_{k-1}\| \\ &\stackrel{(b)}{\leq} \sum_{k=1}^l \frac{C}{C_0} \|Z_{k-1}\|_{w(\infty)} \\ &\stackrel{(c)}{\leq} \frac{C}{C_0} \left(1 + \sum_{k=2}^l \frac{1}{\sqrt{\log n}} \left(\frac{1}{2}\right)^{k-1} \right) \|Z_0\|_{w(\infty)} \\ &\leq \frac{2C}{C_0} \|Z_0\|_{w(\infty)} \\ &\stackrel{(d)}{\leq} \frac{1}{8}, \end{aligned}$$

where (a) follows because $Z_{k-1} \in T$, (b) follows from (27), (c) follows from (28) and (29), and (d) follows from (25) and C_0 is sufficiently large.

Furthermore, by applying the spectral norm bound on random matrix in [19], we have

$$\|\mathcal{P}_{T^\perp}(\lambda \operatorname{sgn}(S))\| \leq \lambda \|\operatorname{sgn}(S)\| \leq \lambda \cdot 4\sqrt{n}. \quad (30)$$

Since $\lambda = \frac{1}{32\sqrt{n \log n}}$, we have

$$\|\mathcal{P}_{T^\perp}(\lambda \operatorname{sgn}(S))\| \leq \frac{1}{8\sqrt{\log n}} \leq \frac{1}{8}.$$

In order to show that Y satisfies (10), we derive

$$\begin{aligned} \|Y\|_\infty &= \left\| \sum_{k=1}^l \mathcal{R}_{\Gamma_k} Z_{k-1} \right\|_\infty \\ &\stackrel{(a)}{\leq} \left\| \sum_{i,j} \frac{6}{p_{ij}} \mathbb{I}_{\{(i,j) \in \Gamma_1\}} (Z_0)_{ij} e_i e_j^* \right\|_\infty + \sum_{k=2}^l \left\| \sum_{i,j} \frac{1}{q_{ij}} \mathbb{I}_{\{(i,j) \in \Gamma_k\}} (Z_{k-1})_{ij} e_i e_j^* \right\|_\infty \\ &\leq 6 \cdot \max_{i,j} \frac{|(Z_0)_{ij}|}{p_{ij}} + \sum_{k=2}^l \max_{i,j} \frac{|(Z_{k-1})_{ij}|}{q_{ij}} \\ &\leq \frac{6}{C_0 \sqrt{n \log n}} \|Z_0\|_{w(\infty)} + \sum_{k=2}^l \frac{1}{C_0 \sqrt{n}} \|Z_{k-1}\|_{w(\infty)} \\ &\stackrel{(b)}{\leq} \frac{6}{C_0 \sqrt{n \log n}} \|Z_0\|_{w(\infty)} + \sum_{k=2}^l \frac{1}{C_0 \sqrt{n \log n}} \left(\frac{1}{2}\right)^{k-1} \|Z_0\|_{w(\infty)} \\ &\leq \frac{7}{C_0 \sqrt{n \log n}} \|Z_0\|_{w(\infty)} \\ &\stackrel{(c)}{\leq} \frac{224C}{C_0} \lambda \\ &\stackrel{(d)}{\leq} \frac{\lambda}{4}, \end{aligned}$$

where (a) is due to the golfing scheme with non-uniform partitions, (b) follows from (28) and (29), (c) follows from (25) and (d) follows because C_0 is sufficiently large.

B Proofs of Key Properties

In this section, we prove the key lemmas provided in Appendix A.1. The central technique used here is non-communicative Bernstein inequality [20].

B.1 Proof of Lemma 2

We note that the condition $\|\mathcal{P}_T - \mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T\| \leq \frac{1}{2}$ implies for any matrix Z

$$\frac{1}{2} \|\mathcal{P}_T Z\|_F \leq \|\mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T(Z)\|_F \leq \frac{3}{2} \|\mathcal{P}_T Z\|_F.$$

Thus, for any matrix Z , we have

$$\begin{aligned} \left\| \mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T(Z) \right\|_F^2 &= \langle \mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T(Z), \mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T(Z) \rangle \\ &= \langle Z, (\mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T)^* \mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T(Z) \rangle \\ &= \langle \mathcal{P}_T(Z), \mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T(Z) \rangle \\ &\leq \|\mathcal{P}_T Z\|_F \|\mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T(Z)\|_F \\ &\leq \frac{3}{2} \|\mathcal{P}_T Z\|_F^2. \end{aligned}$$

Thus, $\left\| \mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T \right\| \leq \sqrt{3/2}$ and hence $\left\| \mathcal{P}_T \mathcal{R}_{\Omega_0}^{1/2} \right\| \leq \sqrt{3/2}$ because $\mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T$ and $\mathcal{P}_T \mathcal{R}_{\Omega_0}^{1/2}$ are adjoint operators and have equal norm. On the other hand, we show $\left\| \mathcal{R}_{\Omega_0}^{1/2} \right\| \leq 1/\sqrt{p_0}$ as follows. For any matrix

Z ,

$$\begin{aligned}\left\|\mathcal{R}_{\Omega_0}^{1/2}(Z)\right\|_F^2 &= \left\|\sum_{i,j} \frac{1}{\sqrt{p_{ij}}} \mathbb{I}_{\{(i,j) \in \Omega_0\}} Z_{ij} e_i e_j^*\right\|_F^2 \\ &\leq \sum_{i,j} \frac{Z_{ij}^2}{p_{ij}} \leq \frac{1}{p_0} \|Z\|_F^2.\end{aligned}$$

Thus, $\|\mathcal{R}_{\Omega_0} \mathcal{P}_T\| \leq \|\mathcal{R}_{\Omega_0}^{1/2}\| \cdot \|\mathcal{R}_{\Omega_0}^{1/2} \mathcal{P}_T\| \leq \sqrt{\frac{3}{2p_0}}$. Thus, $\|\mathcal{P}_T \mathcal{R}_{\Omega_0}\| \leq \sqrt{\frac{3}{2p_0}}$.

Since we have $\frac{1}{2} \|\mathcal{P}_T Z\|_F \leq \|\mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T(Z)\|_F \leq \frac{3}{2} \|\mathcal{P}_T Z\|_F$ for any matrix $Z \in T$, the operator $\mathcal{P}_T \mathcal{R}_{\Omega_0} \mathcal{P}_T$ mapping T onto itself is well conditioned. Thus, $\mathcal{P}_{\Omega_0} \mathcal{P}_T$ is injective on T , i.e., for $Z \in T$, $\mathcal{P}_{\Omega_0} \mathcal{P}_T(Z) = 0$ if and only if $Z = 0$.

B.2 Proof of Lemma 4

Let δ_{ij} denote the Bernoulli random variable $\mathbb{I}_{\{(i,j) \in \Omega_0\}}$. We can derive

$$\begin{aligned}(\mathcal{R}_{\Omega_0} - I)Z &= \sum_{i,j} \left(\frac{1}{q_{ij}} \delta_{ij} - 1\right) \langle e_i e_j^*, Z \rangle e_i e_j^* \\ &=: \sum_{i,j} X_{ij}.\end{aligned}$$

We note that X_{ij} for all $i, j \in [n]$ are zero-mean independent random matrices. Furthermore,

$$\|X_{ij}\| \leq \frac{1}{q_{ij}} |Z_{ij}| \leq \frac{1}{C_0 \sqrt{n}} \|Z\|_{w(\infty)}.$$

and

$$\begin{aligned}\left\|\sum_{i,j} \mathbb{E} X_{ij} X_{ij}^*\right\| &= \left\|\sum_{i,j} \mathbb{E} \left(\frac{1}{q_{ij}} \delta_{ij} - 1\right)^2 Z_{ij}^2 e_i e_i^*\right\| \\ &= \left\|\sum_{i,j} \left(\frac{1}{q_{ij}} - 1\right) Z_{ij}^2 e_i e_i^*\right\| \\ &\leq \max_i \sum_j \frac{Z_{ij}^2}{q_{ij}} \\ &\leq n \|Z\|_{w(\infty)}^2 \cdot \max_{i,j} \frac{w_{ij}^2}{q_{ij}} \\ &\leq \|Z\|_{w(\infty)}^2 \cdot \frac{1}{C_0^2} \max_{i,j} \left(C_0 \sqrt{\frac{\mu_{ij} \Gamma}{n}}\right) \\ &\leq \frac{1}{C_0^2 \log n} \|Z\|_{w(\infty)}^2\end{aligned}$$

Similarly, it can be shown that $\|\sum_{i,j} \mathbb{E} X_{ij}^* X_{ij}\| \leq \frac{1}{C_0^2 \log n} \|Z\|_{w(\infty)}^2$. Thus, applying the non-commutative Bernstein inequality, we obtain

$$\begin{aligned}\|(\mathcal{R}_{\Omega_0} - I)Z\| &= \left\|\sum_{i,j} X_{ij}\right\| \\ &\leq C \left(\sqrt{\frac{1}{C_0^2 \log n} \|Z\|_{w(\infty)}^2 \cdot \log n} + \frac{1}{C_0 \sqrt{n}} \|Z\|_{w(\infty)} \cdot \log n\right) \\ &\leq \frac{C}{C_0} \|Z\|_{w(\infty)}\end{aligned}$$

with high probability.

B.3 Proof of Lemma 5

For any entry index pair (a, b) , we have

$$\begin{aligned}
& [(\mathcal{P}_T \mathcal{R}_{\Omega_0} - \mathcal{P}_T)Z]_{ab} \cdot \frac{1}{w_{ab}} \\
&= \sum_{i,j} \left(\frac{1}{q_{ij}} \delta_{ij} - 1 \right) Z_{ij} \langle \mathcal{P}_T(e_i e_j^*), e_a e_b^* \rangle \cdot \frac{1}{w_{ab}} \\
&= \sum_{i,j} \left(\frac{1}{q_{ij}} \delta_{ij} - 1 \right) Z_{ij} \langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle \cdot \frac{1}{w_{ab}} \\
&=: \sum_{i,j} x_{ij}.
\end{aligned}$$

We note that x_{ij} for $i, j \in [n]$ are independent random variables and $\mathbb{E}x_{ij} = 0$. Furthermore,

$$\begin{aligned}
|x_{ij}| &\leq \frac{1}{q_{ij}} |Z_{ij}| \cdot |\langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle| \cdot \frac{1}{w_{ab}} \\
&\leq |Z_{ij}| \cdot \frac{1}{C_0 \beta^{-2} \sqrt{\mu_{ij} r/n}} \cdot \sqrt{\frac{2\mu_{ij} r}{n}} \cdot \sqrt{\frac{2\mu_{ab} r}{n}} \cdot \frac{1}{\sqrt{\frac{\mu_{ab} r}{n^2}}} \\
&\leq \frac{2\beta^2}{C_0} \sqrt{\frac{\mu r}{n}} \frac{|Z_{ij}|}{w_{ij}} \\
&\leq \frac{2\beta^2}{C_0^2 \log n} \|Z\|_{w(\infty)},
\end{aligned}$$

and

$$\begin{aligned}
\left| \sum_{i,j} \mathbb{E}x_{ij}^2 \right| &\leq \sum_{i,j} \mathbb{E} \left(\frac{1}{q_{ij}} \delta_{ij} - 1 \right)^2 Z_{ij}^2 \cdot |\langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle|^2 \cdot \frac{1}{w_{ab}^2} \\
&\leq \sum_{i,j} \left(\frac{1}{q_{ij}} - 1 \right) \frac{Z_{ij}^2}{w_{ij}^2} \cdot \frac{w_{ij}^2}{w_{ab}^2} \cdot |\langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle|^2 \\
&\leq \frac{1}{C_0 \beta^{-2} \sqrt{n\mu r}} \cdot \|Z\|_{w(\infty)}^2 \cdot \frac{1}{\mu_{ab} r} \|\mathcal{P}_T(e_a e_b^*)\|_F^2 \\
&\leq \frac{2\beta^2}{C_0} \sqrt{\frac{\mu r}{n}} \cdot \|Z\|_{w(\infty)}^2 \\
&\leq \frac{2\beta^2}{C_0^2 \log n} \|Z\|_{w(\infty)}^2,
\end{aligned}$$

where we use the fact $\|\mathcal{P}_T(e_a e_b^*)\|_F^2 \leq \frac{2\mu_{ab} r}{n}$, and the last steps of the above two derivations are due to the fact $C_0 \sqrt{\mu r/n} \log n \leq 1$ implied by our assumption.

Thus, applying the non-commutative Bernstein inequality, we have

$$\begin{aligned}
\left| \sum_{i,j} x_{ij} \right| &\leq C \left(\sqrt{\frac{2\beta^2}{C_0^2 \log n} \|Z\|_{w(\infty)}^2 \cdot \log n} + \frac{2\beta^2}{C_0^2 \log n} \|Z\|_{w(\infty)} \cdot \log n \right) \\
&= C \left(\frac{\sqrt{2}}{C_0} \beta + \frac{2}{C_0^2} \beta^2 \right) \|Z\|_{w(\infty)} \\
&\leq \frac{1}{2} \beta \|Z\|_{w(\infty)},
\end{aligned}$$

with high probability, provided that C_0 is sufficiently large.

B.4 Proof of Lemma 6

We first express $(\mathcal{P}_T \operatorname{sgn}(S))_{ab}$ as

$$\begin{aligned}
\langle e_a e_b^*, \mathcal{P}_T \operatorname{sgn}(S) \rangle &= \langle \operatorname{sgn}(S), \mathcal{P}_T(e_a e_b^*) \rangle \\
&= \sum_{i,j} \delta_{ij} \langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle \\
&=: \sum_{i,j} x_{ij}
\end{aligned}$$

where

$$\delta_{ij} = \begin{cases} 1 & \text{with prob. } \rho_{ij}/2 \\ 0 & \text{with prob. } 1 - \rho_{ij} \\ -1 & \text{with prob. } \rho_{ij}/2. \end{cases}$$

We note that x_{ij} for $i, j \in [n]$ are independent random variables and $\mathbb{E}x_{ij} = 0$. Furthermore, by applying Cauchy-Schwartz inequality and the fact $\|\mathcal{P}_T(e_a e_b^*)\|_F^2 \leq \frac{2\mu_{ab}r}{n}$, we have

$$|x_{ij}| \leq |\langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle| \leq \sqrt{\frac{2\mu r}{n}} \cdot \sqrt{\frac{2\mu_{ab}r}{n}}$$

and

$$\begin{aligned} \left| \sum_{i,j} \mathbb{E}x_{ij}^2 \right| &= \left| \sum_{i,j} \mathbb{E}\delta_{ij}^2 \langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle^2 \right| \\ &= \left| \sum_{i,j} \rho_{ij} \langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle^2 \right| \\ &\leq \left| \sum_{i,j} \langle e_i e_j^*, \mathcal{P}_T(e_a e_b^*) \rangle^2 \right| \\ &= \|\mathcal{P}_T(e_a e_b^*)\|_F^2 \\ &\leq \frac{2\mu_{ab}r}{n}. \end{aligned}$$

Thus, applying the non-commutative Bernstein inequality, we obtain

$$\begin{aligned} \left| \sum_{i,j} x_{ij} \right| &\leq C \left(\sqrt{\frac{2\mu_{ab}r}{n} \cdot \log n} + \sqrt{\frac{2\mu r}{n}} \cdot \sqrt{\frac{2\mu_{ab}r}{n} \cdot \log n} \right) \\ &\leq C \sqrt{\frac{\mu_{ab}r \log n}{n}}, \end{aligned}$$

where the last inequality follows from the fact $C_0 \sqrt{\mu r/n} \log n \leq 1$ implied by the assumption.